

# Copula theory and applications: Part I

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# Outline

Introduction: Background

Some important definitions and results

Distribution function

Quantile function

Copula: definition

Some history

Some references

## Example 1

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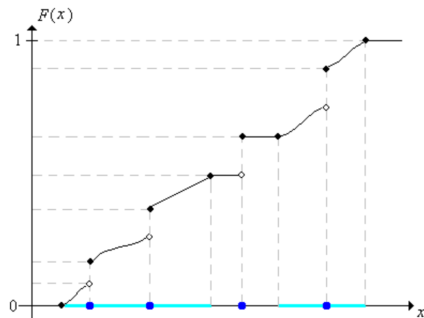
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COPULAS ALLOW TO MODEL MULTIVARIATE  
(STOCHASTIC) DEPENDENCY SEPARATELY FROM THE  
MARGINALS

## Some important definitions

1.  $F$  is non-decreasing:  $x_1 < x_2 \implies F(x_1) \leq F(x_2)$ ;
  2.  $F$  is right-continuous:  $\lim_{h \rightarrow 0^+} F(x + h) = F(x)$ ;
  3.  $\lim_{x \rightarrow -\infty} F(x) = 0$  and  $\lim_{x \rightarrow \infty} F(x) = 1$ ;
  4.  $F$  has at most a countable number of discontinuities.
- If a function  $F : \mathbb{R} \rightarrow [0, 1]$  is non-decreasing, right-continuous,  $\lim_{x \rightarrow -\infty} F(x) = 0$ , and  $\lim_{x \rightarrow \infty} F(x) = 1$ , then  $F$  is the **distribution function** of some random variable.

# Distribution function



## Some important definitions

A distribution function  $F$  is called

- ▶ **discrete** if for some countable set of numbers  $\{x_j\}$  and point masses  $\{p_j\}$ ,

$$F(x) = \sum_{x_j \leq x} p_j, \text{ for all } x \in \mathbb{R}.$$

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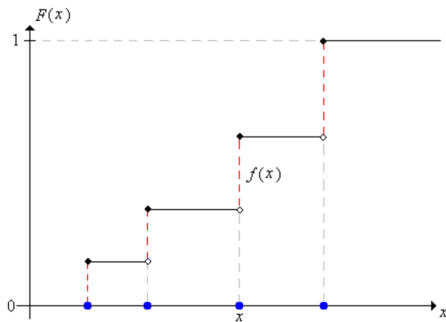
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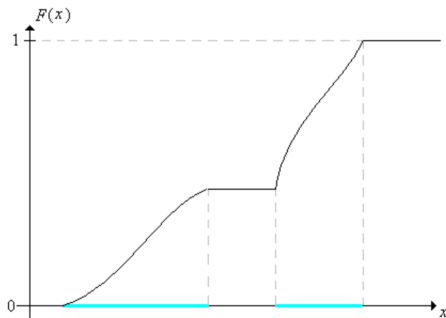
The function  $f$  is called **density** of  $F$ ;

- ▶ **singular** if  $F \neq 0$ ,  $F'$  exists and equals 0 a.e. (almost everywhere).

# Discrete distribution function



# Continuous distribution function



## Some important definitions

The **joint, or multivariate, distribution function** of random vector  $\mathbf{X} = (X_1, X_2, \dots, X_n)'$

$$F_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = P(X_1 \leq x_1, X_2 \leq x_2, \dots, X_n \leq x_n),$$

for  $x_k \in \mathbb{R}$ ,  $k = 1, 2, \dots, n$ . This is written more compactly as

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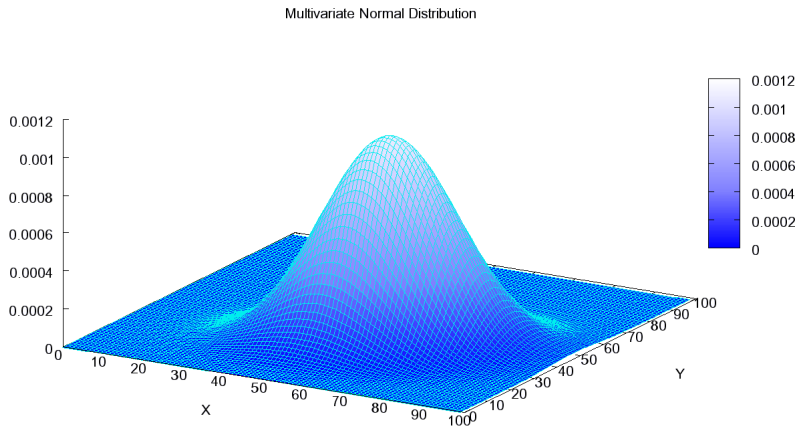
- ▶ For discrete distributions the **joint probability mass function** is defined by

$$p_{\mathbf{X}}(\mathbf{x}) = P(\mathbf{X} = \mathbf{x}), \quad \mathbf{x} \in \mathbb{R}^n$$

- ▶ and in the absolutely continuous case we have a **joint density**

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{\partial^n F_{\mathbf{X}}(\mathbf{x})}{\partial x_1 \partial x_2 \cdots \partial x_n}, \quad \mathbf{x} \in \mathbb{R}^n.$$

# Bivariate normal density function



## Some important definitions

Let  $(X_1, X_2)$  be a 2-dimensional continuous random vector with joint density function  $f_{12}(x_1, x_2)$ . Then the **marginal density function** is defined as

$$f_1(x_1) = \int_{-\infty}^{\infty} f_{12}(x_1, x_2) dx_2.$$

and the **marginal distribution function** is defined as

$$F_1(x_1) = P(X_1 \leq x_1, X_2 < \infty) = \int_{-\infty}^{x_1} \int_{-\infty}^{\infty} f_{12}(x'_1, x_2) dx'_1 dx_2.$$

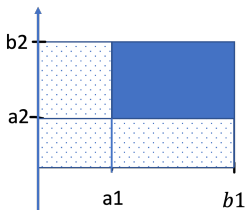


## Proposition

*Necessary and sufficient conditions for a bounded, nondecreasing, and right-continuous function  $F$  on  $\mathbb{R}^2$  to be a bivariate distribution function are:*

1.  $\lim_{x_j \rightarrow -\infty} F(x_1, x_2) = 0, j = 1, 2;$
2.  $\lim_{(x_1, x_2) \rightarrow (\infty, \infty)} F(x_1, x_2) = 1;$
3. (rectangle inequality or 2-increasing fct) *for any  $(a_1, a_2), (b_1, b_2)$  with  $a_1 < b_1, a_2 < b_2,$*

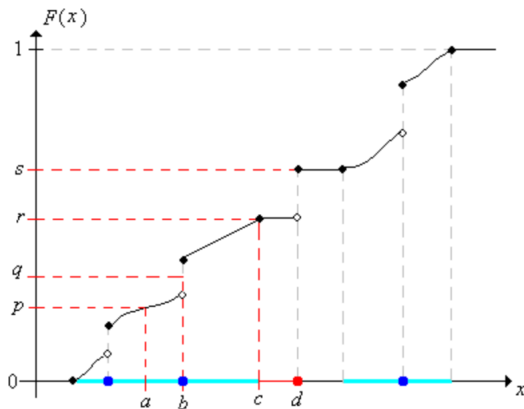
$$F(b_1, b_2) - F(a_1, b_2) - F(b_1, a_2) + F(a_1, a_2) \geq 0.$$



# Quantile function

Let  $X$  be a real-valued random variable with distribution function  $F(x)$ . Then the **quantile function** of  $X$  is defined as

$$Q(p) = F^{-1}(p) = \inf\{x : F(x) \geq p\}, \quad 0 \leq p \leq 1. \quad (1)$$



# Quantile function

Note: If  $F(x)$  is continuous and strictly increasing,  $Q(p)$  is the unique value  $x$  such that  $F(x) = p$ .

**Exercise 1** Let  $X$  be an  $\text{Exp}(\lambda)$  random variable:

$$P(X \leq x) = F(x) = 1 - \exp[-\lambda x]$$

for  $x \geq 0$  and  $\lambda > 0$ . Compute the quantile function  $Q(p)$ .

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Solution:

$1 - p = P(X \geq x) = P(1/X \leq 1/x)$ ; thus,  $Q(1 - p) = 1/x$ .

# Uniform distribution on $[a, b]$

For  $a, b \in \mathbb{R}$ ,  $a < b$ , define the **uniform distribution** on  $[a, b]$  by density

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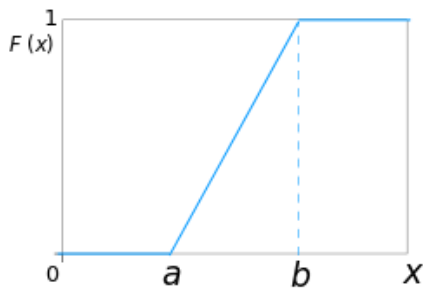
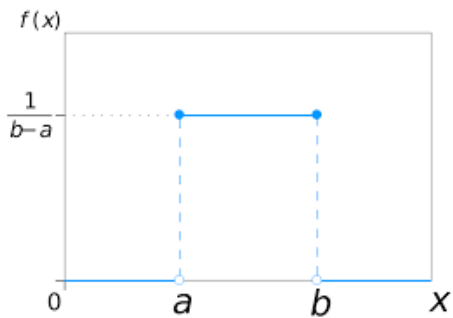
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Let  $U$  be a uniform random variable on  $[0, 1]$ ; then for  $0 \leq c < d \leq 1$ ,

$$P(c \leq U \leq d) = d - c.$$



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$$P(F_X(X) \leq u) = P(X \leq F_X^{-1}(u)) = F_X(F_X^{-1}(u)) = u$$

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Proof:

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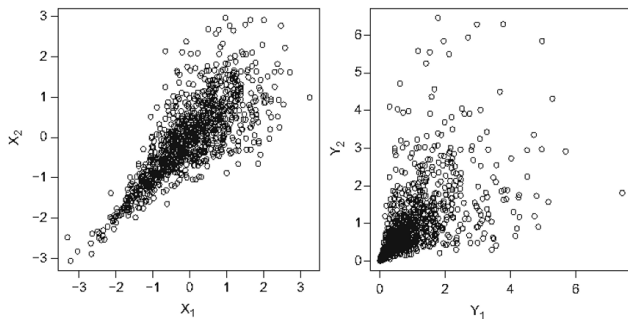
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Proof:  $P(F^{-1}(U) \leq x) = P(U \leq F(x)) = F(x)$ ,  $x \in \mathbb{R}$ .  
 (“quantile transform”)

# A Motivating Example (Hofert, 2018)



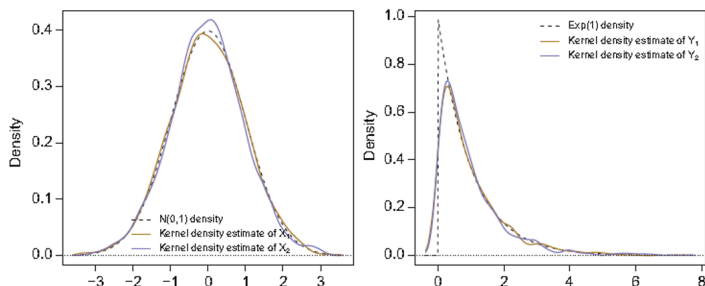
**Fig. 1.1** Scatter plots of  $n = 1000$  independent observations of  $(X_1, X_2)$  (left) and of  $(Y_1, Y_2)$  (right)

$$\text{corr}(X_1, X_2) = .77 \text{ and } \text{corr}(Y_1, Y_2) = .56$$



# A Motivating Example (Hofert, 2018)

The marginal distributions of  $(X_1, X_2)$  and  $(Y_1, Y_2)$

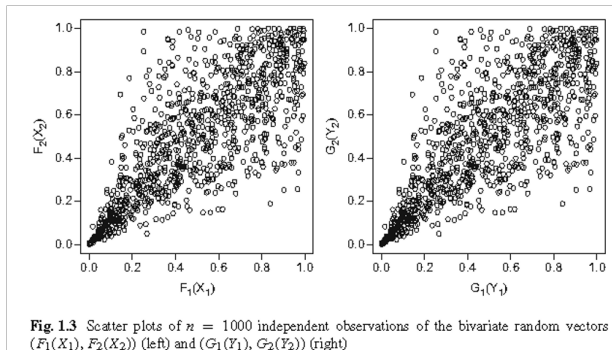


**Fig. 1.2** Kernel density estimates of the densities of  $X_1, X_2$  (left) and  $Y_1, Y_2$  (right). The dashed curves represent the  $N(0, 1)$  (left) and  $\text{Exp}(1)$  (right) densities

If we could transform the two data sets so that they become similar in terms of the underlying marginal dfs, their comparison in terms of dependence would be made on much fairer grounds.

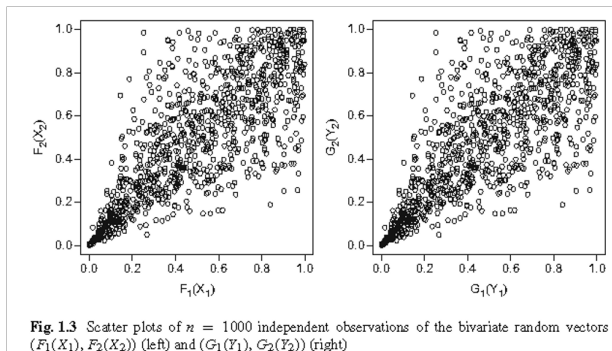
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After a probability transformation of the marginals of  $(X_1, X_2)$  and  $(Y_1, Y_2) \rightarrow (F_1(X_1), F_2(X_2))$  and  $(G_1(Y_1), G_2(Y_2))$



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If dependence between the components of a random vector should not be affected by its marginal distributions, the conclusion is that **the two data sets are indistinguishable in terms of dependence and only differ in terms of the underlying marginal dfs.**

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$$u_2 v_2 - u_2 v_1 - u_1 v_2 + u_1 v_1 = u_2(v_2 - v_1) - u_1(v_2 - v_1) = (u_2 - u_1)(v_2 - v_1) \geq 0$$

## Sklar's Theorem, 1959

(Remember: assuming continuous random variables).

For **any** bivariate distribution function  $F(x_1, x_2)$  with margins  $F_1(x_1)$  and  $F_2(x_2)$  there exists a **unique** copula  $C$  such that

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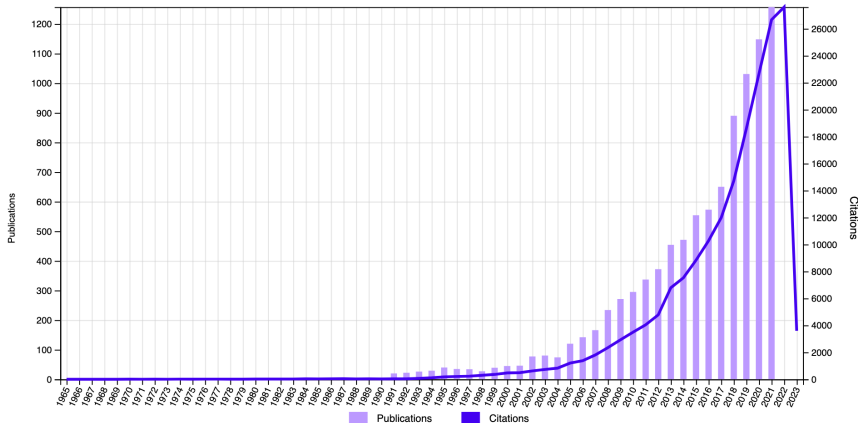
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A copula “couples” the bivariate distribution with its marginals.

- (i) study the structure of stochastic dependency in a “scale-free” manner, i.e., independent of the specific marginal distributions,
- (ii) construct families of multivariate distributions with specified properties.

# Copula publications/citations in Web-of-Science 1995-2021



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## Interview Article Special Issue in memory of Abe Sklar

## Open Access

Christian Genest\*

## A tribute to Abe Sklar

<https://doi.org/10.1515/demo-2021-0110>

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This paper gives an account of the life and works of the American mathematician Abe Sklar. Born in Chicago on November 17, 1925, Sklar completed his PhD at the California Institute of Technology in 1956. He then joined the Illinois Institute of Technology, where he taught mathematics until his retirement in 1995. With his close friend and lifelong collaborator Berthold Schweizer (1929–2010), he was a pioneer of the theory of probabilistic metric spaces, which were introduced in 1942 by the Austro-American mathematician Karl Menger (1902–85). Together, Schweizer and Sklar made important contributions to the algebra of functions, the study of  $t$ -norms, and distributional chaos. Sklar is also credited for the notion of copula and for showing that any multivariate distribution function can be expressed in terms of its univariate margins and a copula. This result, known as Sklar's representation theorem, is the bedrock of a widespread data analytical technique called copula modeling. Sklar passed away in Chicago on October 30, 2020.

By now, just about anyone who is conscious of the role of dependence in data analysis has heard of copulas as a powerful and flexible tool for modeling association and assessing its impact on inference, decision making, and risk management. This approach is rooted in a 3-page note, written in French, which appeared in 1959 in the *Publications de l'Institut de statistique de l'Université de Paris*. This paper [S6], attributed only to “M. Sklar” (M. for Mr.), without address or affiliation, claimed without proof that given any  $d$ -variate cumulative distribution function  $H$  with one-dimensional margins  $F_1, \dots, F_d$ , a function  $C : [0, 1]^d \rightarrow [0, 1]$  having specific analytical properties can always be found such that

$$H = C(F_1, \dots, F_d). \quad (1)$$



FELIX SALMON

BUSINESS FEB 23, 2009 12:00 PM

## Recipe for Disaster: The Formula That Killed Wall Street

In the mid-'80s, Wall Street turned to the quants—brainy financial engineers—to invent new ways to boost profits. Their methods for minting money worked brilliantly... until one of them devastated the global economy.



In the mid-'80s, Wall Street turned to the quants—brainy financial engineers—to invent new ways to boost profits. Their methods for minting money worked brilliantly... until one of them devastated the global economy. \* JIM KRANTZ / INDEX STOCK IMAGERY, INC. / GALLERY STOCK

# The formula that killed Wall Street

$$\Pr[T_A < 1, T_B < 1] = \Phi_2(\Phi^{-1}(F_A(1)), \Phi^{-1}(F_B(1)), \gamma)$$

The formula that killed so many pension plans: David X. Li's Gaussian copula, as first published in 2000. Investors exploited it as a quick – and fatally flawed – way to assess risk.

## Probability

Specifically, this is a joint default probability – the likelihood that any two members of the pool (A and B) will both default. It's what investors are looking for, and the rest of the formula provides the answer.

## Survival times

The amount of time between now and when A and B can be expected to default. Li took the idea from a concept in actuarial science that charts what happens to someone's life expectancy when their spouse dies.

## Equality

A dangerously precise concept, since it leaves no room for error. Clean equations help both quants and their managers forget that the real world contains a surprising amount of uncertainty, fuzziness, and precariousness.

## Copula

This couples (hence the Latin term copula) the individual probabilities associated with A and B to come up with a single number. Errors here massively increase the risk of the whole equation blowing up.

## Distribution functions

The probabilities of how long A and B are likely to survive. Since these are not certainties, they can be dangerous: Small miscalculations may leave you facing much more risk than the formula indicates.

## Gamma

The all-powerful correlation parameter, which reduces correlation to a single constant – something that should be highly improbable, if not impossible. This is the magic number that made Li's copula function irresistible.

Felix Salmon (Wired, 2009; reprinted in Significance, 2012)

## Some references

- ▶ Hofert, M, Kojadinovic, I, Mächler, M., Yan, J. (2018) *Elements of Copula Modeling with R*. Springer-Verlag, UseR! Series
- ▶ Durante, F., & Sempi, C. (2015). *Principles of copula theory*. Boca Raton, FL: CRC Press.
- ▶ Joe, H. (2014). *Dependence modeling with copulas*. Boca Raton, FL: Chapman & Hall/CRC.
- ▶ Jaworski, P., Durante, F., Härdle, W. K., & Rychlik, T. (Eds.) (2010). *Copula theory and its applications*. Lecture notes in statistics (Vol. 198). Berlin: Springer
- ▶ Sklar, A. (1959). *Fonctions de répartition à  $n$  dimensions et leurs marges*. Publications de l'Institut de Statistique de l'Université de Paris, 8, 229–231.
- ▶ Nelsen, R. B. (2006). *An introduction to copulas*. New York: Springer.
- ▶ Groesser & Okhrin (2020) Wires Computational Statistics: *Copulae: Overview and Recent Developments*.

# Some references

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